

Multivariate indices as estimates of dry body weight for comparative study of body size in Lepidoptera

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Abstract. Comparative studies on the size of adult Lepidoptera (moths and butterflies) frequently rely on single linear estimates of body size, namely of forewing length or wingspan. As the shape of the wings of these insects – in fact, of all body parts – differs from one taxon to another, such estimates of body mass may not be adequate for comparisons across a wide taxonomic range. Using the length and width of the forewing, thorax and abdomen, as well as the wing area of 375 species and their correlations with dry body weight, several composite indices were determined that might be used in different circumstances. As the coefficients of determination from the multivariate regression models were rather high ($R^2 > 0.96$), the results are believed to be reliable. A critical re-evaluation of the results indicates that important variations in the regression slopes described here would be expected, if at all, only from species with unusual body shapes. Incidentally, the bivariate relationships are in agreement with former comparative work on Lepidoptera and other terrestrial insects in that the relationship between body weight and single linear measurements follows a slightly negatively allometric trend, implying comparatively lighter bodies at the largest body sizes and relatively heavier ones at the shortest body sizes.

Introduction

As one of the hyper-diverse insect taxa, the order Lepidoptera is well suited for comparative work on subjects of broad biological relevance such as the evolution of body size and its correlation with other traits (e.g., Nilsson and Forsman 2003; Simonsen and Kristensen 2003; Allen et al. 2011; Ribeiro and Freitas 2011; Symonds et al. 2012). This requires an estimate of body size that is valid across distantly related subtaxa, as a broad taxonomic coverage would be of interest for recovering long-term evolutionary trends or patterns.

Although body mass, or weight, is generally accepted as an accurate measure of size for Lepidoptera (e.g., Miller 1977), adult body weight has been rarely used in comparisons across species, and if so, only within a relatively narrow taxonomic framework (e.g., Agosta and Janzen 2005; Davis et al. 2012). In fact, the published data on body weight cover a small number of the known moth and butterfly species. This is largely due to the practical difficulties of obtaining live (fresh) adults from a wide array of taxa and geographic regions for weighing in standard conditions. Most often, the adult size of these insects has been estimated in one of two ways, depending on the purposes of the study. The first consists of using body length or an alternative linear measure (such as head width) to estimate body mass, based on the generally good correlations between those measurements and fresh or dry body weight across large numbers of species of invertebrates (Sample

et al. 1993; Hóðar 1996 and references therein). This approach is frequently utilized in ecological studies on e.g. biomass production or on the diet of insectivore vertebrates (Hóðar 1997; Heyman and Gunnarson 2011; Legagneux et al. 2012) as well as in fresh water ecology (Benke et al. 1999). The second context is that of ecological or evolutionary work on the Lepidoptera based on interspecific comparisons of one linear measurement of the adult wings (generally well correlated to adult body weight: Nylin et al. 1993; Miller 1977, 1997). Here, the most popular metrics are wingspan (the distance between the tips of the forewings of a set specimen, or twice the distance between the tip of one of the forewings to the center of the thorax) and forewing length (e.g., Hawkins and Lawton 1995; Beck and Kitching 2007; Hambäck et al. 2007).

Wings are the most relevant structure of these insects to the human eye, and there are good reasons for wing size to be correlated with body mass for functional reasons, as Lepidoptera are flying insects. However, some degree of structural variation affecting the relationship between wing size and body weight has been documented at several taxonomic levels including the intra-specific one (Van Dyck et al. 1997; Tiple et al. 2009; Shreeve et al. 2009; Symonds et al. 2012). As already stated by Miller (1977), the broad body architecture is likely to differ markedly between the members of distantly related taxa of similar body weights, so that more precise estimates of body mass of species in varied taxonomic positions require a more elaborate combination of linear measurements. It is conceivable that a multivariate approach based on several variables correlated with body weight might achieve this purpose.

The main objective of this study was to determine a composite index based on several linear estimates that could predict accurately the dry body weight of set specimens (e.g., from museum collections or even scale illustrations) irrespective of the species phylogenetic position. The reason for selecting dry body mass instead of fresh body weight is of a practical nature: because these insects are usually preserved as dried samples in scientific collections, the possibility to test and re-elaborate any results is far more feasible than obtaining reliable fresh (live) weights from the same set of species. The second objective was to determine the sensitivity of such an index to sample size (the number of species), taxonomic diversity and morphological heterogeneity as a means to measure its robustness (if it is to be applied to species different from those used to fit it).

Methods

To avoid heterogeneity caused by the patterns of sexual dimorphism in adult size, the comparison was restricted to adult males from any available source, totaling 665 individuals from 375 species distributed among 61 families. The selection emphasized the diversity of size within and across families and included samples from any region in the world that could be processed.

Measurements

The measurements were performed on dry set (pinned or spread), complete male specimens. When fresh adults were available, these were first dried in the position traditionally used for these insects in entomological collections. The measures described below were taken in one of four ways: (a) under a stereomicroscope with an ocular micrometer, (b) on a digitized scale drawing made with an optical camera lucida adapted to a stereomicroscope ($\times 10$ to $\times 40$), (c) on a digital photograph of the specimen taken together with a standard scale bar, taken either

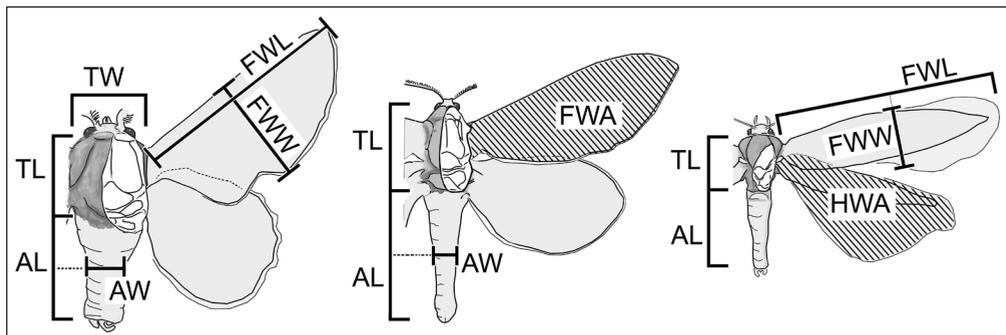


Figure 1. Slightly idealized representations of three typical adult Lepidoptera (left to right: Lasiocampidae, Hepialidae, Gelechiidae) to illustrate the variables measured. The right side of the thoraces is represented as devoid of the scale cover to make more evident the limits of this tagma. The three drawings are scaled to the same forewing length. Linear measurements are indicated by bars and areas by a striped pattern. FWL = fore wing length, FWW = forewing width, FWA = forewing area, HWA = hind wing area, TL = thorax length, TW = thorax width, AL = abdomen length, AW = abdomen width.

with a macro lens (up to 1:1) or on a photo microscope at low magnification, or (d) with a Vernier caliper (exceptionally in the case of some of the largest moths). The program ImageJ (Rasband 2012) was used to measure the digitized images.

Six linear measurements (in mm) were taken (Figure 1): thorax length (TL), thorax width (TW, taking the point of insertion of the fore wings as a reference), abdomen length (AL) excluding terminal hair pencils or protruding genital appendages, abdomen width (AW, taken at the midpoint of the line represented by AL), forewing length (FWL, from the insertion of the wing on its costal margin to its apex including the fimbriae) and forewing width (FWW, the distance between edges following a line perpendicular to FWL at its midpoint). In addition, the area of the fore- and hindwings (including the fringes) were recorded (FWA, HWA, as mm²). The mean species values are available as Supplementary material (Suppl. material 1: nexus format text).

Repeated measures and replicates

To estimate the magnitude of error measurement, the mean within sample and mean within species coefficients of variation were calculated after replicated measurements taken on each individual and between individuals within species.

- (1) Every measurement was taken twice for each specimen using two different methods among those detailed above (most frequently a, b and c), on two different dates.
- (2) Whenever possible two male specimens of approximately the same size (judged from wing-span by naked eye) of the species were processed. However, replications were not always possible as data from single representatives of a number of species were included if this contributed to an increase in the taxonomic or geographic coverage of the species selection.

Dry body weight

The insects were dried to a constant weight at 60° for 48 hours (72 h for the largest specimens). The pins, if present, were removed carefully (but see below). The weight of the whole speci-

men was determined to the nearest 0.01 mg in a Mettler AT261 balance (species of wingspan of ca. 15 mm or above) or in a Mettler Toledo XP6 microbalance with precision of 0.001 mg (individuals smaller than that size).

Pinned specimens

Although medium or larger sized collection specimens can generally be de-pinned and remounted without much difficulty, there is always some risk of damage. For a small number of loaned specimens (ca. 20 individuals) the weight of the pins was estimated, then subtracted from that of the dry mounted specimen. Samples of 10 individual pins from four different brands and numbers (gauges): 000, 00, 0 and 1 to 6 (all with nylon heads and 37 mm long) were measured and weighed. The weights were taken to the nearest 0.01 mg, and the widths measured with a precision of 0.0179 mm under a binocular microscope with an ocular scale line. The relationship between the log-transformed weights and widths was highly consistent: $\log_{10}(\text{pin weight in mg}) = 2.339 + 1.908 \log_{10}(\text{pin diameter in mm})$, $R = 0.997$, $P < 0.0001$, $n = 350$.

Small moths

The smallest moths (broadly corresponding to the heterogeneous assemblage of the “microlepidoptera”) posed some special difficulties, which handicapped the use of reference collections as sources of size data. These moths are fragile and very likely to be damaged if treated in the way described above, and even though they are frequently mounted on smaller pins (‘minutiae’, weighting 0.69–3.15 mg for widths of 0.10 and 0.20 mm respectively) the small variation in the length of these tiny metal pieces represents an excessive error in terms of the specimen dry weight. Moreover, as the genital pieces are of interest for identification, collection specimens frequently lack the abdomen or a large part of it as it was removed for identification. Finally, most of them cannot be easily identified to species level without expertise. For these reasons the data from several families in this category were obtained from a small reference collection at the author’s department. This hosts expert-identified specimens collected two decades ago at a single site, so new samples were taken at the same location during 2011–2012 to reasonably cover the lower part of the size range, although at the cost of low geographic variation.

Multivariate models

All the variables were transformed to their decimal logarithms. This facilitated comparisons with results from earlier research (as most size-weight relations have been modelled using the equation $\text{weight} = a \times \text{size}^b$: Reiss 1989; Ganihar 1997), linear-regression approaches as well as some demands of the comparative method adopted (described below). After log-transformation, all the variables fitted reasonably to the normal distribution with Kolmogorov-Smirnov test values of $d < 0.049$, $P > 0.05$ in all instances (Suppl. material 2: frequency distribution graph).

The multivariate models were fitted using the General Regression Models module of Statistica (Statsoft 2004). For model selection, a manual iterative forward-backwards procedure was adopted to exclude redundant variables.

Independent contrasts and phylogenetic hypothesis

The method of phylogenetically independent contrasts (Felsenstein 1985; Harvey and Pagel 1991) was used to control for phylogenetic effects. The contrasts were calculated using the

software PDAP:PDTREE (Midford et al. 2009) integrated in the package Mesquite (Maddison and Maddison 2011). Branch lengths were set to equal length (1.00), and the polytomies were estimated as single contrasts, which were calculated after the original output.

The working hypothesis on phylogenetic relationships was built according to the classification proposed by van Nieukerken et al. (2011), with the relationships above the family level adapted after the tree topologies from Kawahara and Breinholt (2014) complemented by Regier et al. (2009, 2013), Mutanen et al. (2010), Bazinet et al. (2013) and Martijn et al. (2014). Further information was gathered from other recent literature (details available in Suppl. material 3: documentation on phylogeny).

In the absence of any other references, the formal classifications of Fauna Europaea (Karscholt et al. 2013) for the European species and of the Lepindex database (Beccaloni et al. 2013) for other geographic regions was adopted. The tree was assembled manually; preference was given to the most recent results, or to those with the highest statistical support, but keeping any former hypotheses if these have not been contradicted. Thus, except in face of conflicting evidence the formal taxa at the levels of superfamily, family, subfamily and genus were adopted even when their monophyletic status had not been corroborated in all instances. The tree topology and data are available from the Suppl. material 4 and 1 (4: tree topology, 1: tree nexus format). The resulting dendrogram showed high resolution (ca. 77%), which of course is overoptimistic in terms of strictly phylogenetic criteria.

Regressions were done through the origin to estimate the correlations and slopes. After a multivariate regression model was obtained, Least Squares Regression was used to estimate the intercept for the working data set keeping the evolutionary slopes already obtained.

Robustness of the models

The number of species and of supraspecific taxa available for this study was obviously small if compared to the estimated number of existing species in the order Lepidoptera (more than 150,000 species: van Nieukerken et al. 2011). Thus, one further question can be posed – to what extent are the results presented sensitive to the addition of new taxa? The relationship between the errors in the predicted weight data and the diversity in body size, morphology (excluding body weight) and taxonomy were determined. The underlying idea is that any sources of diversity that are positively correlated to large errors in the predictions should denote species' features liable to modify significantly the models obtained.

The error in the predicted dry body weight (DBW) values were measured as the mean of the absolute values of the residuals from the two best fit models (described below) calculated for randomly selected subsets of n species, where $n = 5, 10, 25, 50, 100, 150, 200, 250, 300$ and 350 . Forty replicates were taken at each n plus one more sample consisting of the whole data set. The taxonomic and structural diversities of each of such 401 species samples were estimated using the following attributes:

- (a) Species diversity: the number of species in each sample.
- (b) Variation in dry body weight: the standard deviation of the log-transformed dry body weights.
- (c) Structural variation. This variable was intended to account for structural/anatomical variation as reflected by the measurements taken, irrespective of body weight. To do this,

each of the eight variables were regressed on body weight, one at a time. The residuals of such bivariate regressions were used as the new variables, now linearly independent of body weight. Applying Principal Component Analysis to this set of residuals (Bartlett's Sphericity test $X^2 = 344.24$, $P < 0.001$; KMO index = 0.72) resulted in three components accounting for 66.96% of the variance (respectively 41.51%, 14.59% and 10.86%). The standard deviation in these three components (weighted by the respective contribution of each component) was used as an index of structural (body shape) diversity, linearly independent from dry weight.

- (d) Taxonomic/phylogenetic diversity. This was tentatively estimated in four alternative ways: (1) Number of clades (absolute number of supra-specific nodes). (2) Phylogenetic diversity (PH): the number of clades or nodes represented in the sample minus one, plus the number of species as defined by Faith (1992), with all branches set to 1.00. (3) Relative Phylogenetic Diversity (RPD, the number of clades above the species level divided by the number of species). And (4) Taxonomic Distinctness (Clarke and Warwick 1998; Allen et al. 2009); this was calculated using the software PAST (Hammer et al. 2001) after simplifying the number of taxonomic categories to 10 which included the suborders, superfamilies, families, subfamilies and genera plus five intermediate levels.

As the relationships between the mean residuals and these variables tended to be asymptotic rather than linear, the bivariate and multivariate regressions were performed using Generalized Regression Models and the logarithmic link function.

Results

Size range

The dry body mass of the selected species covered a range of variation of nearly five orders of magnitude, from 0.03 mg to more than 2 g, corresponding to forewing lengths of between 1.8 mm and 110 mm (see Suppl. material 2 and 5; 2: frequency distribution; 5: mean by superfamily). The lightest and smallest species belonged to the genus *Stigmella* (Nepticulidae, with one male weighting 0.034 mg), while two males of the reputedly longest-winged moth, the Erebiidae *Thysannia agrippina* (Cramer, 1776) (see e.g. Kons 1998) had dry weights of 916–1,300 mg and one male of the Saturniidae *Attacus atlas* (L., 1758) weighed 1,126 mg. However the heaviest specimen weighed belonged to the hawk-moth family (*Cocytius* sp., Sphingidae, which exceeded 2.1 grams).

The replicated measurements (Table 1) suggested that the forewing and thoracic linear dimensions may reflect lower proportions of error than the abdomen length or width measurements when taken of the same specimen. Although the estimates between pairs of individuals from the same species differed to some extent, it was clear that the highest amount of variation was accounted for by the abdomen data. Forewing length appeared to be even more constant than the thorax measurements within individuals. This might reflect a bias in the observer's abilities, although it is also likely that the reference landmarks to measure wing length (the tegulae and the tip of the wing) are more obvious than the other reference structures, especially when the body is coated by a dense cover of hair-like scales.

Table 1. Estimate of measurement error for dry body weight and six linear measurements, measured as a percentage of the mean. The values given are the mean coefficients of variation (100·CV) (± 1 SD) averaged across individuals (from duplicated measurements on each specimen, $n = 662$) and from different replicates of the same species (within species, $n = 328$).

| | Within individuals | Within species |
|-----------------------|--------------------|--------------------|
| Dry weight (DBW) | ---- | 13.334 \pm 9.905 |
| Forewing length (FWL) | 2.317 \pm 2.477 | 5.706 \pm 4.138 |
| Forewing width (FWW) | 3.177 \pm 3.843 | 6.174 \pm 6.826 |
| Thorax length (TL) | 3.760 \pm 3.915 | 5.611 \pm 4.748 |
| Thorax width (TW) | 3.032 \pm 3.345 | 5.424 \pm 4.901 |
| Abdomen length (AL) | 4.450 \pm 4.499 | 8.631 \pm 6.769 |
| Abdomen width (AW) | 5.982 \pm 6.473 | 9.541 \pm 6.678 |

Table 2. Relationships between dry body weight and the test variables based on the species mean values, estimated both by bivariate regression (left four columns) and in a multivariate regression model (right three columns; intercept = -0.489, multiple $R = 0.983$, adjusted $R^2 = 0.965$). The β values represent the relative contribution of each variable in the multivariate model.

| Variable | Bivariate regression | | | | Multivariate regression | | |
|----------|----------------------|-------|--------|-----------|-------------------------|--------|--------|
| | R | Slope | P | Intercept | β | Slope | P |
| FWL | 0.939 | 2.772 | <0.001 | -2.137 | -0.060 | -0.178 | 0.359 |
| FWW | 0.920 | 1.989 | <0.001 | -0.320 | -0.044 | -0.095 | 0.390 |
| TL | 0.975 | 2.718 | <0.001 | -0.445 | 0.407 | 1.135 | <0.001 |
| TW | 0.957 | 2.902 | <0.001 | -0.173 | 0.189 | 0.572 | <0.001 |
| AL | 0.948 | 2.790 | <0.001 | -1.173 | 0.082 | 0.241 | 0.029 |
| AW | 0.936 | 2.529 | <0.001 | 0.553 | 0.150 | 0.404 | <0.001 |
| FWA | 0.941 | 1.266 | <0.001 | -1.174 | 0.274 | 0.368 | 0.008 |
| HWA | 0.926 | 1.279 | <0.001 | -1.136 | 0.011 | 0.015 | 0.862 |

Bivariate regressions and preliminary multivariate regressions

The results from bivariate regressions of DBW on the other variables as well as the full multivariate results (with all the variables in the model) are presented in Table 2 (species means, all $R > 0.92$) and Table 3 (independent contrasts, all $R > 0.82$). The effects of the linear estimates of wing size (FWL and FWW), although significant in the bivariate comparisons performed on the species data, were outweighed by those of the forewing area (FWA) in the multivariate approach. Across the contrasts, FWL had a significant but negative effect in the regression models suggesting a complex relationship between body weight and wing size and shape.

Multivariate regression model selection

Several alternative models fit by stepwise regression were calculated with multiple R values above 0.979 in all instances. Models 1 and 2 (Table 4; Figure 2) are those with the highest multivariate R based in the species raw data and in the independent contrasts respectively. These two models included the effects of wing area, which may be more difficult to measure in spread specimens. However, because of their highest fits they were used as the basis for the last/next

Table 3. Relationships between dry body weight and the test variables based on the independent contrasts, estimated by bivariate regression (left three columns) and by multivariate regression (right three columns; multiple $R = 0.914$, adjusted multiple $R^2 = 0.833$). All regressions were forced through the origin (no intercept). The β values represent the relative contribution of each variable in the multivariate model.

| Variable | Bivariate regression | | | Multivariate regression | | |
|----------|----------------------|-------|--------|-------------------------|--------|--------|
| | R | Slope | P | β | Slope | P |
| FWL | 0.835 | 2.489 | <0.001 | -0.146 | -0.434 | 0.091 |
| FWW | 0.813 | 2.132 | <0.001 | 0.040 | 0.104 | 0.547 |
| TL | 0.891 | 2.663 | <0.001 | 0.376 | 1.122 | <0.001 |
| TW | 0.859 | 2.632 | <0.001 | 0.185 | 0.568 | 0.001 |
| AL | 0.817 | 2.353 | <0.001 | 0.055 | 0.159 | 0.257 |
| AW | 0.817 | 2.185 | <0.001 | 0.149 | 0.398 | 0.003 |
| FWA | 0.840 | 1.153 | <0.001 | 0.301 | 0.448 | 0.003 |
| HWA | 0.821 | 1.210 | <0.001 | 0.015 | 0.022 | 0.843 |

Table 4. The two multivariate models with highest R scores among those fitted using the species mean values (1) and the phylogenetically independent contrasts (2). The statistics given are the coefficients of the intercepts and slopes (Coeff.), β values (relative contribution of each variable after standardization) and P (significance). The multivariate statistics are represented at the base of the table. The regression based on the independent contrasts was done through the origin (without intercept, statistics in the two bottom rows); the intercept given (-0.553) was fitted a posteriori for the species values in the data set using the slopes (coefficients) stated.

| | (1) Species means | | | (2) Independent Contrasts | | |
|------------------|-------------------------------------|---------|--------|-------------------------------------|---------|--------|
| | Coeff. | β | P | Coeff. | β | P |
| Intercept | -0.180 | --- | 0.207 | -0.553 | --- | <0.001 |
| FWL | -0.745 | -0.252 | 0.015 | --- | --- | --- |
| FWL ² | 0.183 | 0.148 | 0.013 | --- | --- | --- |
| FWA | 0.346 | 0.257 | <0.001 | --- | --- | --- |
| TL | 1.149 | 0.412 | <0.001 | 1.087 | 0.395 | <0.001 |
| TW | 0.622 | 0.205 | <0.001 | 0.616 | 0.167 | <0.001 |
| AL | 0.312 | 0.106 | 0.005 | --- | --- | --- |
| AW | 0.368 | 0.136 | <0.001 | 0.408 | 0.109 | <0.001 |
| FWA | --- | --- | --- | 0.378 | 0.294 | <0.001 |
| Model statistics | | | | | | |
| R | 0.9828 | | | 0.981 | | |
| $F(P)$ | $F_{7, 367} = 1489.83 (P < 0.0001)$ | | | $F_{4, 371} = 1409.32 (P < 0.0001)$ | | |
| R [origin] | --- | | | 0.9140 | | |
| $F(P)$ [origin] | --- | | | $F_{3, 287} = 351.54 (P < 0.0001)$ | | |

step. Several alternatives (Suppl. material 6: alternative models) should allow estimations of DBW in circumstances that are frequent in entomological collections such as specimens without abdomen or with its distal end missing due to identifications based in the external genitalia.

Robustness of the models

The regressions of the estimated error of the predictions (measured as the mean of the absolute value of the residuals) on the indicators of taxonomic, size and structural diversity led to the same

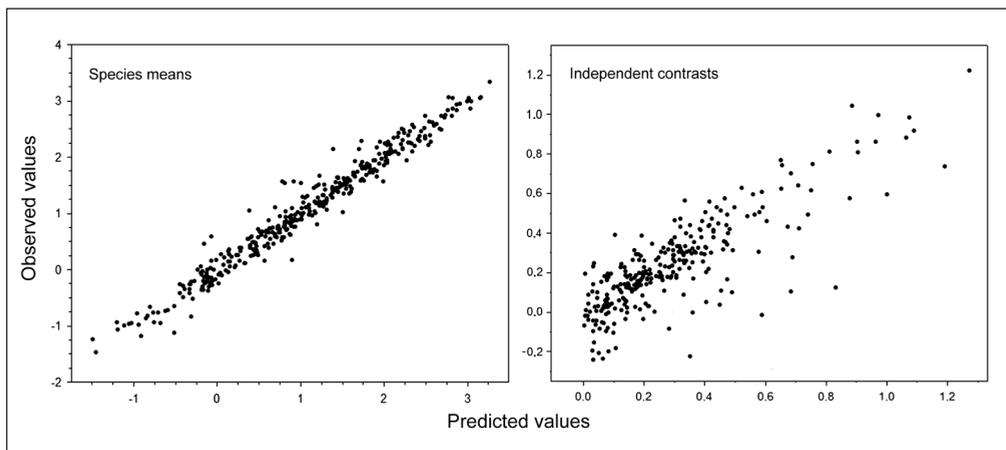


Figure 2. Dispersion plots illustrating the fit (predicted on observed weights) of the two multivariate models of highest R^2 scores based on the raw species data (above) and the independent contrasts (below) (respectively, models 1 and 2 in Table 4).

Table 5. Sensitivity of the best models to several sources of diversity in the species selected. Relationships between the deviations of the predicted data (mean absolute residuals from 401 subsets of 5–375 species) based on the multivariate models 1 and 2 (from Table 4) and several alternative estimates of structural diversity (number of species, taxonomic and phylogenetic diversity, morphology and body weight), estimated through multiple regression. The contributions of the variables are represented in the upper (Coeff. = coefficient, Wald = Wald’s statistic) and the multivariate statistics in the lower rows. The Ordinary Least Squares (OLS) R^2 values calculated a posteriori for the two multiple regression models are given for comparison. PH = Phylogenetic diversity, RPD = Relative Phylogenetic Diversity.

| Variable | Model 1 | | | Model 2 | | |
|-------------------------|--------------------|--------|---------|--------------------|--------|---------|
| | Coeff. | Wald | P | Coeff. | Wald | P |
| Number of species | 0.0003 | 1.837 | 0.175 | 0.0003 | 1.166 | 0.280 |
| Body Weight diversity | 0.0125 | 1.752 | 0.186 | 0.0053 | 0.268 | 0.604 |
| Morphological diversity | 0.0965 | 40.349 | <0.0001 | 0.0867 | 27.582 | <0.0001 |
| Taxonomic distinctness | 0.0032 | 0.718 | 0.396 | 0.0018 | 0.195 | 0.659 |
| Number of clades | -0.0003 | 1.917 | 0.166 | -0.0002 | 1.191 | 0.275 |
| PH | -0.00002 | 0.014 | 0.906 | -0.00003 | 0.027 | 0.870 |
| RPD | -0.0143 | 16.371 | <0.0001 | -0.0161 | 17.527 | <0.0001 |
| Model statistics | | | | | | |
| Deviance/DF | 0.0022 | | | 0.0033 | | |
| Log-likelihood | 470.817 | | | 445.012 | | |
| OLS R^2 (P) | 0.168 (P < 0.0001) | | | 0.163 (P < 0.0001) | | |

results in the bivariate and multiple tests, irrespective of the data analyzed (species values or independent contrasts); thus, for simplicity, only the multivariate results are presented in Table 5. Only two of the variables had significant effects with opposite signs: morphological diversity (with a positive coefficient) and the relative phylogenetic diversity (with a negative effect).

Discussion

The results generally show high correlations between all linear dimensions of the Lepidopteran body, or the wing areas, and total dry body weight. This is not surprising given the relatively important range of sizes covered and, especially, because a functional link between the variables measured and total body size should exist in insects that must be able to fly effectively such as the male specimens of moth and butterfly species studied.

The results are consistent with the fact that the wings of Lepidoptera are thin structures (thus relatively light even if comparatively broad and evident) while the largest proportion of the body weight is determined by the weight of the main thoracic and abdominal structures. Forewing length is a popular estimate of body size in butterflies and moths as it is easier to measure than other body dimensions. However, this measure has by itself a lower predictive power of dry body weight than the thoracic dimensions (length and width) or, depending on the method used, abdomen length. Thus, wingspan, taken as the distance from the midpoint of the thorax to the tip of the forewing, would in theory be more accurate than the length of the wing alone as it would partly account for thorax width. However, as stated by Miller (1977) the estimate of 'wingspan' most widely used in the specialized literature is the distance between the tips of the two forewings, where the spreading technique is a potential source of error. Alternatively, some of the body dimensions, especially the abdomen width, tend to be measured with lower accuracy than wing size. In spread collection specimens, the abdomen is frequently deformed and contracted to different degrees, and measurements made on the thorax may be hindered by the dense scale/hair clothing of some of these insects. Under these circumstances a composed 'body size index' appears to be a practical alternative measurement to body weight, particularly when different species are to be compared.

For the linear measurements that are more directly related to body length, such as the thoracic and abdominal lengths, the slopes determined across the species means (2.7–2.8, see Table 2) are exactly in the same range as those found for the relationship between body length and dry mass in terrestrial and aquatic insects on a wider taxonomic scope (2.6 to 2.9: Rogers et al. 1976; Schoenert 1980; Bugherr and Meyer 1997; Benke et al. 1999), or within the order Lepidoptera (Ganihar 1997). Hódar (1996) obtained slopes in the range 2.8–2.9 for the regressions of body weight on head width for butterflies and moths. This supports the idea that dry body mass correlates to single linear measurements such as body length following a slightly negative allometric trend (that is, with a slope slightly below 3.0 which would be expected for the volume to length ratio), at least if estimated by Least Squares Regression. Values of the slope based on the independent contrasts tend to be more conservative (Table 3). However generalizing on these grounds remains difficult since single linear surrogates of body weight may well vary among taxa (e.g. from 2.1 to 2.9 between two families of Lepidoptera; Miller 1977, 1997).

Among the several drawbacks of the present results is the fact that intraspecific variation has not been controlled for, and cannot be distinguished from other sources of error. This may be acceptable under the assumption that intraspecific variation in body weight is generally higher than interspecific variation for the same trait. Given this and the widespread phenomenon that intraspecific allometric trends follow different (generally less steep) slopes than the interspecific trends in animal taxa (e.g. Harvey and Pagel 1991), one corollary is that the body mass indexes

presented here are probably not suitable for determining dry body weights accurately within a species. One further limitation of the results presented concerns the estimation of dry body weight in living or fresh (not dried) adults of Lepidoptera, because all the body parts experience some degree of contraction after drying (including the wings; Van Hook et al. 2012); these effects are especially noticeable in the abdomen. In such cases, a suboptimal model (Suppl. material 6: alternative models) could be used as an approximation, or alternatively the bivariate relationships of body weight to forewing length or area as given in Table 2.

Of course, it is likely that the predictive accuracy of the regression models selected can be improved by spreading the selection of species. The results in Table 5 suggest that this would neither be achieved simply by increasing the number of species compared nor by broadening their variance in body weight; instead, it seems that the amount of error in the predictions is primarily correlated with the proportion of morphological diversity of the species compared (irrespective of their body weight) relative to their phylogenetic diversity. In other words, the results may be relatively stable unless for species selections featured by extreme variations in wing and body shape, from subtaxa of Lepidoptera not represented in the sample analyzed.

Although the comparative method of independent contrasts is statistically robust in the absence of accurate estimates of branch lengths, the contrasts are calculated by dividing the differences between each pair of values at a node by the estimated evolutionary distances (derived directly from the branch lengths; Felsenstein 1985). This is a source of uncertainty when the precise value of the regression slopes is of interest. Further, the overall value for the slope of a relationship within a large taxon may represent, in some instances, the average of several slopes featuring the different subtaxa (e.g., for butterflies: García-Barros 2002). Thus, although the formulae derived from the independent contrasts might be suitable for the estimation of dry body weight in species from taxa not prospected in this work, it may be subject to criticism and re-evaluation. The fact that their fit to the data was slightly lower than that based on the raw species data may simply reflect some degree of over-sampling on closely related species, but on the basis of the results and for species similar to those selected preference is given to model 1 (Table 4), or alternatively to models 5 and 6 (presented in Suppl. material 6: alternative models).

Conclusion

The fact that the multivariate approaches presented here showed high R^2 scores (> 0.94) for a much wider range of size, morphology and taxonomic variety than that in any former comparable study on Lepidoptera suggest that, although liable to be refined, they may represent a useful tool for comparative work when a wide taxonomic scope is necessary.

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Supplementary material 1

Nexus format text.

Authors: Enrique Gracia-Barros

Data type: Adobe PDF file

Explanation note: Tree topology for the phylogenetic hypothesis adopted, to be used as input in applications reading nexus (requires some slight previous edition).

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Supplementary material 2

Frequency distribution graph.

Authors: Enrique García-Barros

Data type: Adobe TIF file

Explanation note: Frequency distribution of the dry body weight data (mg) across the species studied.

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Supplementary material 3

Documentation on phylogeny.

Authors: Enrique García-Barros

Data type: Adobe PDF file

Explanation note: This is a list of references including the most relevant sources of information used to build the hypothesis on phylogenetic relationships which were not quoted in the main text.

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Supplementary material 4

Tree topology.

Authors: Enrique García-Barros

Data type: Adobe PDF file

Explanation note: Graphic display (dendrogram) to show the hypothesis on phylogenetic relations adopted in this work, after the sources quoted in the main text and in the file: Supplementary material 3.

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Supplementary material 5

Mean by superfamily.

Authors: Enrique García-Barros

Data type: Adobe PDF file

Explanation note: Mean dry body weight and wing length by superfamily, and sample sizes.

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Supplementary material 6

Alternative models.

Authors: Enrique García-Barros

Data type: Adobe PDF file

Explanation note: Alternative or suboptimal regression models derived from the species means or from the independent contrasts.

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